

Estimation of the parameters of a diffusion with discontinuous coefficients

Antoine Lejay & Paolo Pigato

TOSCA, Inria Nancy-Grand Est

PDE & Probability Methods for Interactions

Sophia-Antipolis, March 2017

The Oscillating Brownian motion (OBM)

Terminology of Keilson & Wellner (1978)

$$X_t = x + \int_0^t \sigma(X_s) dB_s, \quad \sigma(x) = \begin{cases} \sigma_+ & \text{if } x \geq 0 \\ \sigma_- & \text{if } x < 0. \end{cases}$$

- Strong existence, uniqueness (Le Gall, 1978)
- Analytic formula of the density, occupation time (Keilson & Wellner, 1978)
- Convergence of the Euler scheme (Chan & Stramer, 1989, Yan, 2002, ...)
- Approximation by Random Walks (Keilson & Wellner, 1978; Helland 1982, ...)

The Oscillating Brownian motion (OBM)

Terminology of Keilson & Wellner (1978)

$$X_t = x + \int_0^t \sigma(X_s) dB_s, \quad \sigma(x) = \begin{cases} \sigma_+ & \text{if } x \geq 0 \\ \sigma_- & \text{if } x < 0. \end{cases}$$

Applications

- In a continuous time version of the Self-Exciting Threshold Auto-Regressive (SETAR) models (Tong, 1983).
- In finance, it mimicks a **leverage effect** of log-prices.

The realized volatility

Observations: $X_{kT/n}$, $k = 0, \dots, n$ (high-frequency data)

How to estimate σ_+ and σ_- ?

Realized volatility type estimators

$$\sigma_{\pm}(n)^2 = \frac{\sum_{k=1}^n (X_{kT/n}^{\pm} - X_{(k-1)T/n}^{\pm})^2}{\frac{T}{n} \sum_{k=1}^n \mathbb{1}_{\pm X_{kT/n} \geq 0}}$$

Idea Itô-Tanaka formula \implies

$$X_t^{\pm} = X_0^{\pm} + \sigma_{\pm} \int_0^t \mathbb{1}_{\pm X_s \geq 0} dB_s + \frac{1}{2} L_t(X)$$

with

- $L_t(X)$ **local time** at 0 (finite variation process).
- $\langle \int_0^{\cdot} \mathbb{1}_{\pm X_s \geq 0} dB_s \rangle_t = Q_t^{\pm}$ **occupation time** of \mathbb{R}_{\pm}

A convergence result

- (i) $\sigma_{\pm}(n)$ is a consistent estimator of σ_{\pm} as $n \rightarrow \infty$.
- (ii) When $T = 1$,

$$\sqrt{n}(\sigma_{\pm}(n)^2 - \sigma_{\pm}^2) \xrightarrow[n \rightarrow \infty]{\text{stable}} \frac{\sqrt{2}\sigma_{\pm}^2}{Q_1^{\pm}} \int_0^1 \mathbb{1}_{\pm X_s \geq 0} d\tilde{B}_s - \frac{1}{Q_1^{\pm}} \frac{2\sqrt{2}}{3\sqrt{\pi}} \frac{\sigma_- \sigma_+}{\sigma_- + \sigma_+} L_1(X),$$

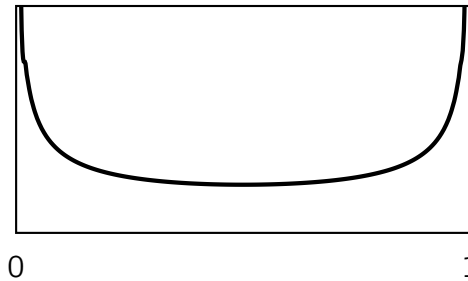
\tilde{B} is a BM indep. from B .

Remarks

- Joint convergence of $(\sigma_+(n), \sigma_-(n))$.
- Using Girsanov's theorem, we could consider the presence of drift (the limit laws are changed).
- By scaling, high-frequency estimation = long time estimation (not true in presence of drift).

Comments

- The limit depends on Q_1^\pm which follows a law of ArcSine type



- \implies either Q_1^+ or Q_1^- is likely to be close to 1
- \implies either σ_+ or σ_- is likely to be loosely estimated.
- The process X is **null recurrent**
 - \implies the limit law is a mixture of normal distribution.
- There is an asymptotic bias which is due to the **discontinuity**.

Some ingredients of the proof

We have to prove, in particular, convergences of type

- $\sqrt{n}[L(B), L(B)] \xrightarrow[n \rightarrow \infty]{\text{proba}} \frac{4\sqrt{2}}{3\sqrt{\pi}} L(B)$
- $\sqrt{n}[L(X), X] \xrightarrow[n \rightarrow \infty]{\text{proba}} 0$
- $\sqrt{n}[L(X), |X|] \xrightarrow[n \rightarrow \infty]{\text{proba}} 0$

with $[Y, Z] = \sum_{i=1}^n (Y_{i/n} - Y_{(i-1)/n})(Z_{i/n} - Z_{(i-1)/n})$

For this, we use that for a suitably decreasing function f ,

$$\frac{1}{\sqrt{n}} \sum_{i=1}^n f(\sqrt{n}X_{i/n}) \xrightarrow[n \rightarrow \infty]{\text{proba}} c(f)L_1(X)$$

by adapting some results of J. Jacod (1998) to the OBM by reducing it to a **Skew Brownian motion** $Y_t = B_t + \gamma L_t(Y)$. Computations are based on explicit expression of the density (this limits immediate generalizations).

Removing the asymptotic bias

Our estimator is changed to

$$\widehat{\sigma}_{\pm}(n)^2 = \frac{\sum_{k=1}^n (X_{k/n}^{\pm} - X_{(k-1)/n}^{\pm})(X_{k/n} - X_{(k-1)/n})}{\frac{1}{n} \sum_{k=1}^n \mathbb{1}_{X_{k/n} \geq 0}}$$

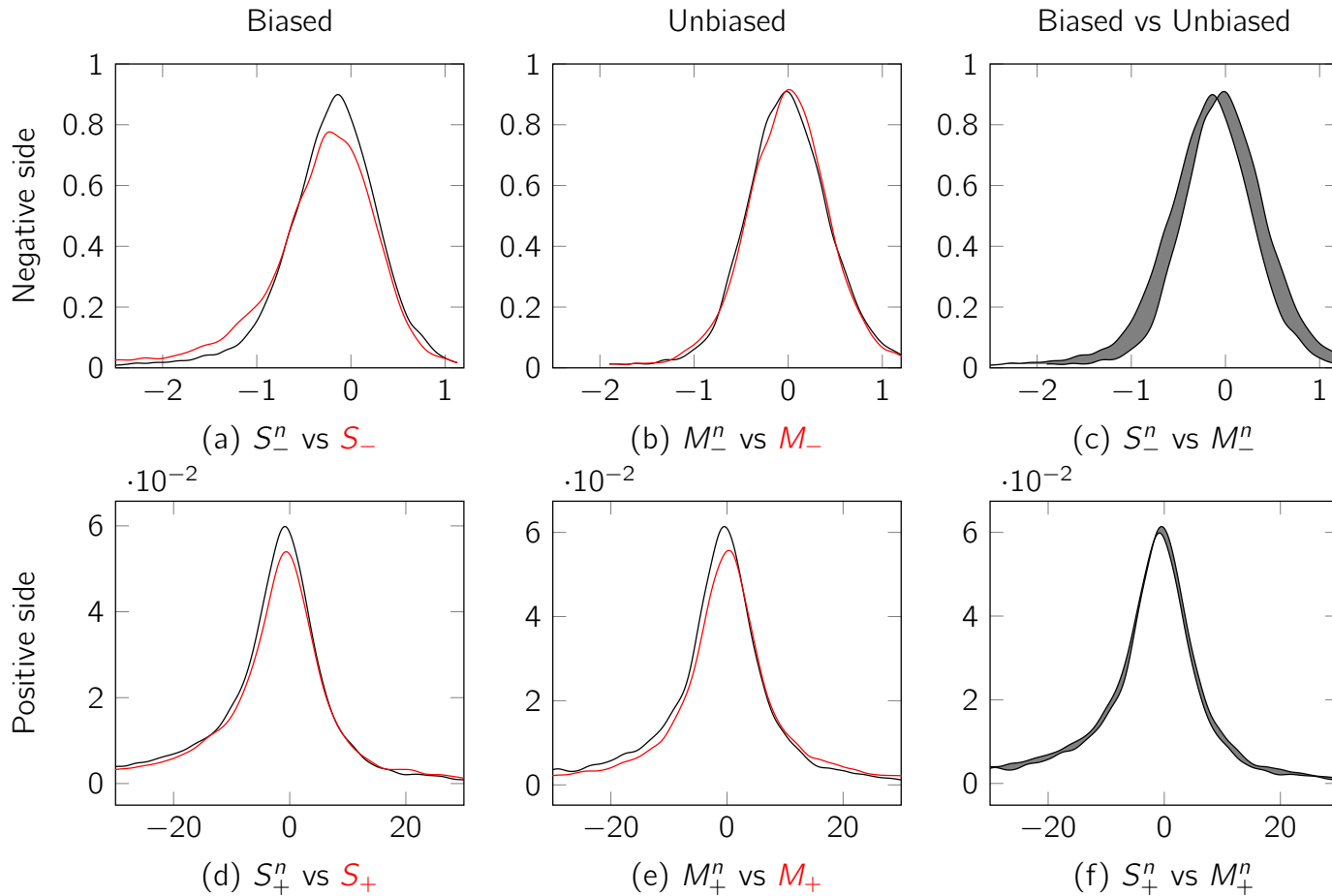
$$\sqrt{n}(\widehat{\sigma}_{\pm}(n)^2 - \sigma_{\pm})^2 \xrightarrow[n \rightarrow \infty]{\text{stable}} \frac{\sqrt{2}\sigma_{\pm}^2}{Q_1^{\pm}} \int_0^1 \mathbb{1}_{\pm X_s \geq 0} d\widetilde{B}_s.$$

The reason is that

$$\sqrt{n} \sum_{k=1}^n (X_{k/n}^+ - X_{(k-1)/n}^+)(X_{k/n}^- - X_{(k-1)/n}^-) \xrightarrow[n \rightarrow \infty]{\text{stable}} \frac{2\sqrt{2}}{3\sqrt{\pi}} \frac{\sigma_- \sigma_+}{\sigma_- + \sigma_+} L_1(X).$$

Numerical illustration on

$$S_+^n := \sqrt{n}(\sigma_+(n) - \sigma)$$

$$M_+^n := \sqrt{n}(\hat{\sigma}_+(n) - \sigma)$$


$\sigma_- = 1/2$, $\sigma_+ = 2$, $n = 500$, on 10 000 paths

Estimation of a two-valued drift

$$X_t = x + \int_0^t \sigma(X_s) dB_s + \int_0^t b(X_s) ds,$$

with

$$\sigma(x) = \begin{cases} \sigma_+ & \text{if } x \geq 0 \\ \sigma_- & \text{if } x < 0. \end{cases} \quad \text{and } b(x) = \begin{cases} b_+ & \text{if } x \geq 0 \\ b_- & \text{if } x < 0. \end{cases}$$

How to estimate (b_-, b_+) ?

- We should consider **long time estimation**.
- The respective signs of b_+ and b_- are fundamentals:

	$b_+ > 0$	$b_+ = 0$	$b_+ < 0$
$b_- > 0$	transient	null recurrent	ergodic
$b_- = 0$	transient	null recurrent	null recurrent
$b_- < 0$	transient	transient	transient

Estimation of a two-valued drift

Itô-Tanaka formula + Maximization of the Girsanov weight

$$\beta_{\pm} = \pm \frac{X_T^{\pm} - X_0^{\pm} - L_T/2}{Q_T^{\pm}} = b_{\pm} + \frac{M_t^{\pm}}{Q_T^{\pm}}$$

where $M^{\pm} = \int_0^t \mathbb{1}_{\pm X_s \geq 0} dB_s$, $\langle M^{\pm} \rangle = Q^{\pm}$.

\implies Empirical estimator for large T

$$b_{\pm} = \pm \frac{X_T^{\pm} - X_0^{\pm} - \hat{L}_T/2}{\hat{Q}_T^{\pm}}$$

where \hat{Q}^{\pm} , \hat{L} are empirical estimators of Q^{\pm} , L .

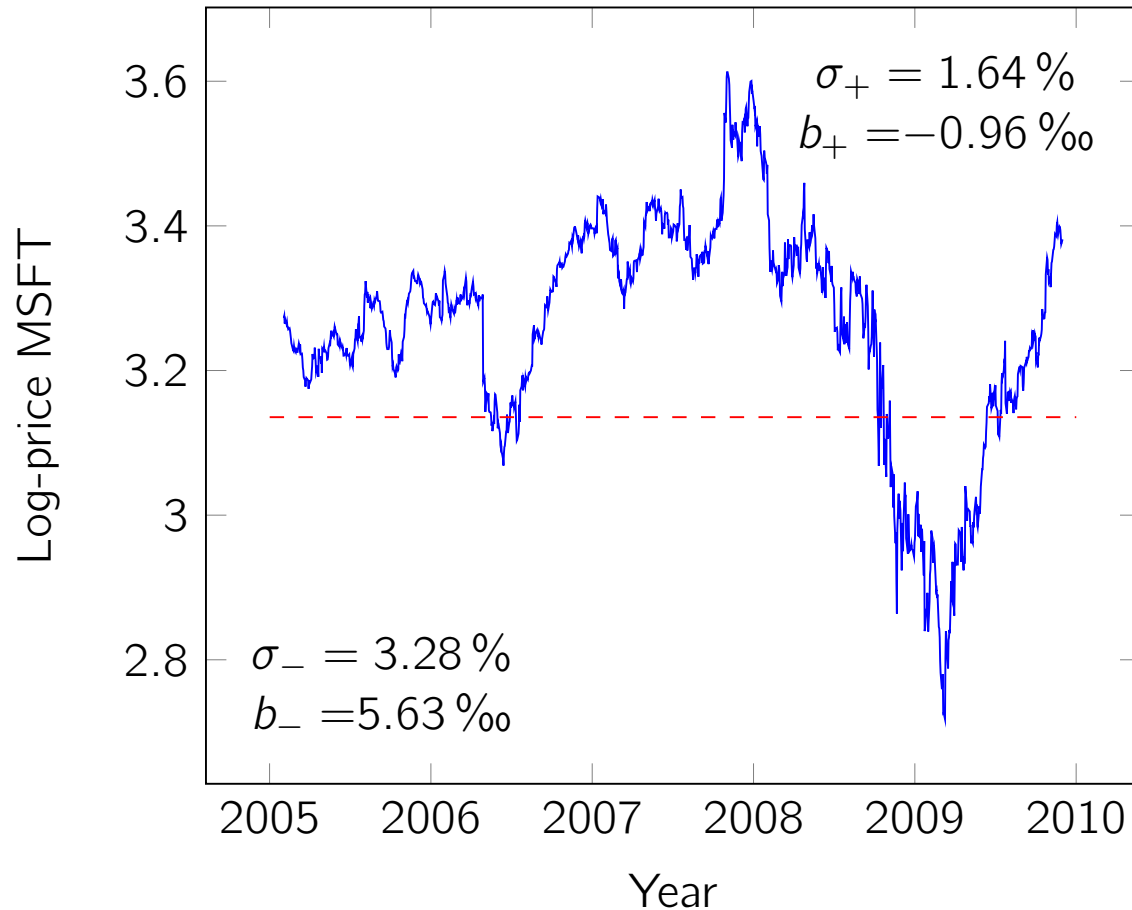
The convergence of the estimator depends on the long time behavior of Q_T^{\pm} , and of the regime of X (work in progress).

Application to financial data

- P. Mota & M. Esquível (2014) have proposed a continuous time version of a SETAR model with delay and threshold regime switching (DTRS).
- Their model uses an artificial thin layer for switching to avoid immediate switchings.
- They propose a least square estimation procedure (coming from time series).
- 21 stocks are analyzed (2005-2010): leverage and mean-reverting effects hold for most of them.

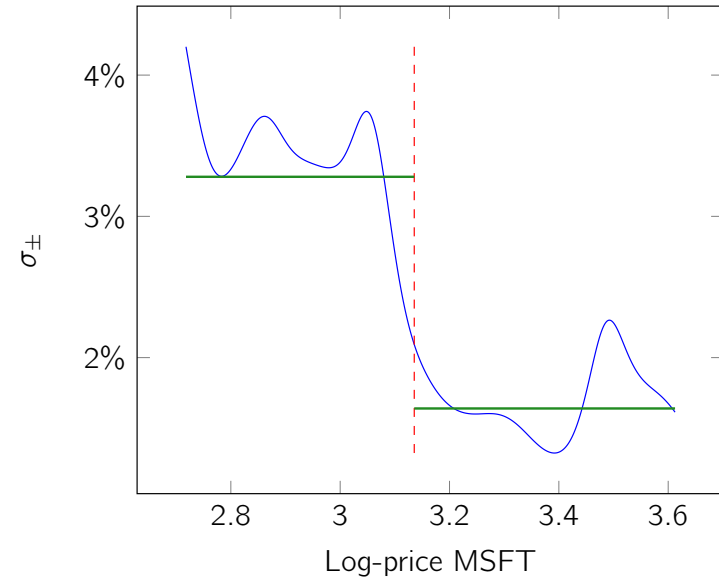
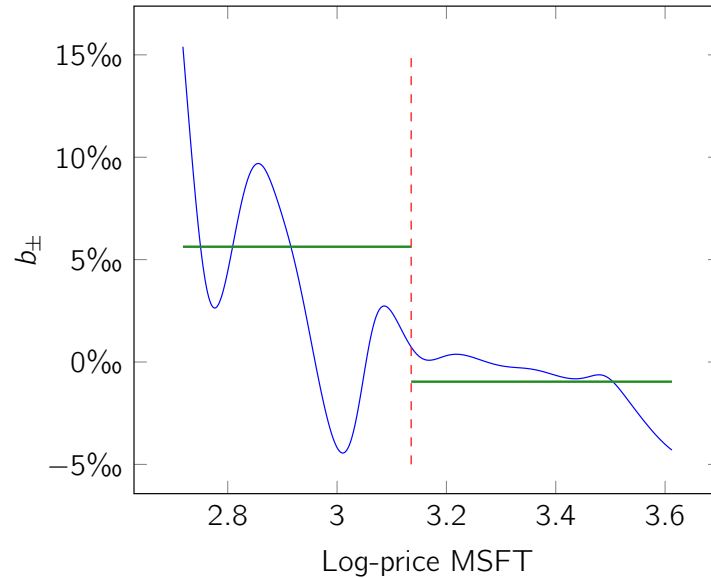
Our estimators gives consistent results with these ones.

Leverage and mean-reverting effects



The threshold is detected with the AIC model selection.

Comparison with a non-parametric estimator



Comparison with a Nadaraya-Watson non-parametric estimation.

Conclusion

- The problem of estimation of SDE with discontinuous coefficients is surprisingly open.
 - Asymptotics of occupation and local times play a very important role.
 - Heavily relies on the limits theorems contained in the book Jacod & Protter. However, they should be adapted to the Skew Brownian motion (some questions are left open).
 - The presence of a drift really change the picture.
-
- ✠ AL & PP, Statistical estimation of the Oscillating Brownian Motion, arxiv:1701.02129 (2017).
 - ✠ Estimation of drift, application to financial data: works in progress.